Recurrence Relation for Single and Product Moments of Record Values from Erlang-truncated Exponential Distribution

Muhammad Mohsin

Department of Mathematics, COMSATS Institute of Information Technology, Lahore, Pakistan

Abstract: Consider X_1 , X_2 , X_3 ,... be a sequence of independently and identically distributed random variables with continuous cumulative distribution function F(x). In this paper, some recurrence relations for single and product moments are derived for Erlang-truncated exponential distribution that are helpful in finding the higher order moments from that of lower order moments.

Key words: Erlang-truncated exponential distribution . upper record values . cumulative distribution function . recurrence relation

INTRODUCTION

A random variable X is said to have Erlangtruncated exponential distribution [1], if its probability density function is of the form

$$f\left(x\right)=\beta\Big(1-e^{-\lambda}\Big)e^{-\beta x\left(1-e^{-\lambda}\right)},0\leq x\leq \infty,\beta>0,\,\lambda>0\ \, (1.1)$$

The distribution function is

$$F(x) = 1 - e^{-\beta x \left(1 - e^{-\lambda}\right)}$$
 (1.2)

Let X_1 , X_2 , X_3 ,... be a sequence of independently and identically distributed random variables with cdf F(x). Set $Y_i = max(min)\{X_1, X_2, ..., X_i\}$ for $i \ge 1$, then X_j is called an upper (lower) record value of $\{X_i, i \ge 1\}$, if

$$Y_j > Y_{j-1}, j > 1, (Y_j < Y_{j-1}, j > 1)$$

It obvious from the definition that X_1 is an upper as well as lower record.

Ahsanullah [2] have also derived the distributional properties of the records by using the Lomax distribution. Balakrishnan [5-7] have obtained the recurrence relations for moments of record values for Gumbel distribution. Some moment properties of the records have been given by Ahsanullah [3, 4]. Nevzorov [8] have given a comprehensive review of the mathematical foundation of the records.

Ahsanullah [4] has given the distribution of k-th upper record; $X_{U(k)}$ as:

$$f_{k:n}(x_k) = \frac{1}{\Gamma(k)} f(x_k) \left[R(x_k) \right]^{k-1}$$
 (1.3)

where

$$R(x) = -\ln\{1 - F(x)\}.$$

The joint distribution of k-th and m-th upper records; $X_{U(k)}$ and $X_{U(m)}$ can be obtained by using the following expression given by Ahsanullah [4]:

$$f_{k,m:n}(x_k, x_m) = \frac{1}{\Gamma(k)\Gamma(m-k-1)}$$

$$r(x_k)f(x_m)[R(x_k)]^{k-1} \qquad (1.4)$$

$$[R(x_m) - R(x_k)]^{m-k-1},$$

with r(x) = R'(x).

In this paper, some recurrence relations for single and product moments are derived for Erlang-truncated exponential distribution in section 2 and 3 respectively. Some concluding remarks are given in section 4.

RECURRENCE RELATION FOR SINGLE MOMENTS

It is easy to note from (1.1) and (1.2) that

$$x f(x) = \left[-\ln(1-F(x))\right](1-F(x))$$
 (2.1)

The relation in (2.1) will be used to establish recurrence relations for moments of the upper record value from Erlang-truncated exponential distribution.

Theorem 2.1: For $n \ge 1$, $k = 0, 1, 2, \dots$ and k < n,

$$E\left(X_{U\left(n+1\right)}^{k}\right) = \left(n-k\right)E\left(X_{U\left(n\right)}^{k}\right)$$

Proof: The kth moment of the n-th upper record is defined as:

$$E\left(X_{U(n)}^{k}\right) = \frac{1}{\Gamma n} \int_{0}^{\infty} x^{k} \left[R\left(x\right)\right]^{n-1} f\left(x\right) dx. \tag{2.2}$$

Using (2.1) in (2.2), we have

$$E\left(X_{U(n)}^{k}\right) = \frac{1}{\Gamma(n)} \int_{0}^{\infty} x^{k-1} \left[-\ln\left(1 - F(x)\right)\right]^{n} \left(1 - F(x)\right) dx. \quad (2.3)$$

Integrating (2.3) by parts taking X^{K-1} as integrating and rest of integrand for differentiation, we obtain:

$$E\left(X_{U\left(n+1\right)}^{k}\right) = \left(n-k\right) E\left(X_{U\left(n\right)}^{k}\right) ; k \le n$$

RECURRENCE RELATIONS FOR PRODUCT MOMENTS

In this section some recurrence relation for product moments for Erlang-truncated exponential distribution have been developed.

Theorem 3.1: For $\le m \le n$, r, s = 0, 1, 2,... and $m \ge (r+1)$,

$$\mathrm{E}\Big(X_{U\left(m+l\right)}^{s+r+1}\Big) = \frac{m - \left(r+1\right)}{m}\; \mathrm{E}\Big(X_{U\left(m\right)}^{r+l}\; X_{U\left(n\right)}^{s}\Big)$$

And

For $n \ge m+2$, $r, s = 0, 1, 2, \dots$ and m > (r+1)

$$E\left(X_{U\left(m\right)}^{r+1}X_{u\left(n\right)}^{s}\right) = \frac{m - \left(r+1\right)}{m} \ E\left(X_{U\left(m+1\right)}^{r+1}\ X_{U\left(n-1\right)}^{s}\right)$$

Proof: We have

$$\begin{split} E\Big(X_{U\left(m\right)}^{r+1}\ X_{U\left(n\right)}^{s}\Big) &= \int\limits_{0}^{\infty}\int\limits_{y}^{\infty}x^{r+1}y^{s}\ f_{m,n}\left(x,y\right)dx\ dy. \\ &= \frac{1}{\Gamma\left(m\right)\Gamma\left(n-m\right)}\int\limits_{0}^{\infty}\int\limits_{y}^{\infty}x^{r+1}y^{s}\Big[-\ln\left(1-f\left(x\right)\right)\Big]^{m-1}\Big[-\ln\left(1-F\left(y\right)\right)+\ln\left(1-F\left(x\right)\right)\Big]^{n-m-1}\frac{f\left(x\right)}{\Big[1-f\left(x\right)\Big]}f\left(y\right)dxdy. \end{aligned} \\ &= \frac{1}{\Gamma\left(m\right)\Gamma\left(n-m\right)}\int\limits_{0}^{\infty}\int\limits_{y}^{\infty}x^{r+1}y^{s}\Big[-\ln\left(1-f\left(x\right)\right)\Big]^{m-1}\Big[-\ln\left(1-F\left(y\right)\right)+\ln\left(1-F\left(x\right)\right)\Big]^{n-m-1}\frac{f\left(x\right)}{\Big[1-f\left(x\right)\Big]}f\left(y\right)dxdy. \end{split}$$

Substituting (2.1) in (3.1) we get

$$E\left(X_{U\left(m\right)}^{r+1}\ X_{U\left(n\right)}^{s}\right) = \frac{1}{\Gamma\left(m\right)\Gamma\left(n-m\right)} \int\limits_{0}^{\infty} \int\limits_{y}^{\infty} x^{r} \ y^{s} \Big[-\ln\left(1-F\left(x\right)\right)\Big]^{m} \Big[-\ln\left(1-F\left(y\right)\right) + \ln\left(1-F\left(x\right)\right)\Big]^{n-m-1} \ f\left(y\right) \ dx \ dy.$$

$$= \frac{1}{\Gamma(m)} \int_{0}^{\infty} y^{s} f(y) I(y) dy.$$
 (3.2)

Where

$$I(y) = \int_{y}^{\infty} x^{r} \left[-\ln(1 - F(x)) \right]^{m} \left[-\ln(1 - F(y)) + \ln(1 - F(x)) \right]^{n-m-1} dx.$$

$$(3.3)$$

Case-1. If n = m + 1.

Integrating (3.3) by parts taking X^{r} as integrating and rest of integrand for differentiation, we obtain:

$$I\left(y\right) = \left[-\left\{-\ln\left(1 - F\left(y\right)\right)\right\}^{m} \frac{y^{r+1}}{r+1}\right] + \frac{m}{r+1} \int\limits_{y}^{\infty} x^{r+1} \left[-\ln\left(1 - F\left(x\right)\right)\right]^{m-1} \frac{f\left(x\right)}{1 - F\left(x\right)} \ dx.$$

Substituting the value of I(y) in (3.2) and simplifying we get

$$\begin{split} E\Big(X_{U(m)}^{r+1}\,X_{U(n)}^{s}\Big) &= \frac{1}{r+1} \begin{bmatrix} \frac{m}{\Gamma(m)} \int\limits_{0}^{\infty} \int\limits_{y}^{\infty} x^{r+1}\,y^{s} \Big[-\ln\left(1-F(x)\right) \Big]^{m-1} \,\frac{f\left(x\right)\,f\left(y\right)}{1-F(x)} \,\,dx\,\,dy \\ \\ -\frac{m}{\Gamma(m+1)} \int\limits_{0}^{\infty} y^{s+r+1} \Big[-\ln\left(1-F(x)\right) \Big]^{m}\,f\left(y\right)dy \\ E\Big(X_{U(m)}^{r+1}\,X_{U(n)}^{s}\Big) &= \frac{m}{(r+1)} \left[\,\,E\Big(X_{U(m)}^{r+1}\,X_{U(n)}^{s}\Big) \,\,-E\Big(X_{U(m+1)}^{s+r+1}\Big) \,\,\Big] \end{split}$$

or

$$E(X_{U(m+1)}^{s+r+1}) = \frac{m - (r+1)}{m} E(X_{U(m)}^{r+1} X_{U(n)}^{s})$$
(3.4)

Case-2: When $n \ge m+2$,

Again, integrating (3.3) by parts taking X^r as integrating and rest of integrand for differentiation, we obtain:

$$\begin{split} I(y) &= \frac{m}{r+1} \int\limits_{y}^{\infty} x^{r+1} \Big[-\ln \big(1 - F\left(x \right) \big) \Big]^{m-1} \Big[-\ln \big(1 - F\left(y \right) \big) + \ln \big(1 - F(x) \big) \Big]^{n-m-1} \\ &\frac{f(x)}{1 - F(x)} \, dx - \frac{\left(n - m - 1 \right)}{r+1} \int\limits_{y}^{\infty} x^{r+1} \left[-\ln \big(1 - f\left(x \right) \big) \right]^{m} \Big[-\ln \big(1 - F(y) \big) + \ln \big(1 - F(x) \big) \Big]^{n-m-2} \frac{f(x)}{1 - F(x)} \, dx. \end{split} \tag{3.5}$$

Putting (3.5) in (3.2), we get

$$E\left(X_{U(m)}^{r+1}X_{U(n)}^{s}\right) = \frac{1}{\Gamma(n)\Gamma(n-m)} \left[\frac{m}{r+1} \int_{0}^{\infty} \int_{y}^{\infty} x^{r+1} y^{s} \left[-\ln(1-F(x))\right]^{m-1} \left[-\ln(1-F(y)) + \ln(1-F(x))\right]^{n-m-1} \right.$$

$$\frac{f(x)}{1-F(x)} f(y) dx dy - \frac{(n-m-1)}{r+1} \int_{0}^{\infty} \int_{y}^{\infty} x^{r+1} y^{s} \left[-\ln(1-F(x))\right]^{m}$$

$$\left[-\ln(1-F(y)) + \ln(1-F(x))\right]^{n-m-2} \frac{f(x)}{1-F(x)} f(y) dx dy.$$
(3.6)

$$\begin{split} & E\left(X_{U(m)}^{r+1}X_{U(n)}^{S}\right) = \frac{m}{(r+1)} E\left(X_{U(m)}^{r+1}X_{U(n)}^{s}\right) - \frac{m}{(r+1)} E\left(X_{U(m+1)}^{r+1}X_{U(n-1)}^{s}\right) \\ & E\left(X_{U(m)}^{r+1}X_{U(n)}^{s}\right) = \frac{m - (r+1)}{m} E\left(X_{U(m+1)}^{r+1}X_{U(n-1)}^{s}\right), \qquad r, s > 0. \end{split}$$

Corollary: From (3.4) substituting the value of $E\left(X_{U(m)}^{r+1}X_{U(n)}^{s}\right)$ in (3.7), we get

$$E\left(X_{U\left(m+1\right)}^{s+r+1}\right) = \left[\frac{m - \left(r+1\right)}{m}\right]^{2} E\left(X_{U\left(m+1\right)}^{r+1} X_{U\left(n-1\right)}^{s}\right)$$
(3.8)

CONCLUSION

The recurrence relations for single and product moments of the record statistics for the Erlang-truncated exponential distribution are derived in this paper. Recurrence relations are useful to characterize the distribution and to reduce the number of operation necessary to obtain a general form for the function under consideration. It can be hoped that following the same procedure researcher may derived the recurrence relation for other continuous distributions.

REFERENCES

- 1. El-Alosey, A.R., 2007. Random sum of New Type of Mixture of distribution, (IJSS), 2 (1): 49-57.
- 2. Ahsanullah, M., 1991. Record values of Lomax distribution, *Statisti*. *Nederlandica*, 41 (1): 21-29.
- 3. Ahsanullah, M., 1992. Record values of independent and identically distributed continuous random variables, Pak. J. Statist., 8 (2): 9-34.

- 4. Ahsanullah, M., 1995. Record Statistics, Nova Science Publishers, USA.
- Balakrishnan, N. and M. Ahsanullah, 1994. Recurrence Relation for single and product moments of record values from exponential generalized Pareto distribution. Comm. Stat. Methods, 23 (10): 2841-2852.
- Balakrishnan, N. and M. Ahsanullah, 1995.
 Relation for single and product moments of record values from exponential distribution. J. Appl. Stat. Sci., 13 (1): 73-80.
- 7. Balakrishnan, N. and M. Ahsanullah, 1994. Recurrence Relations for single and product moments of record values from Lomax distribution, Sankhya, 56 B (2): 140-146.
- 8. Nevzorve, V.B., 2001. Records: Mathematical Theory. American Mathematical Society, Providence, RI.